Introduction
Partial and Overall IEO
Application to the Netherlands
Conclusion

Not All Transitions are Equal

The distribution of education and the contribution of each transition to inequality of educational opportunity

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Outline

Introduction

Partial and Overall IEO

Application to the Netherlands

Generalizing Mare to tracked system

Differences in IEO between cohorts and men and women

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- Mare(1981) showed that overall IEO depended on the transition probabilities, and
- that patrial IEO, as measured by odds ratios, do not.
- As a consequence the relation between partial and over IEO has largely been treated as a black box by:
 - Estimating both partial and overall IEOs and just noticing that they show different trends.
 - ▶ Showing the equation: $\sum_{k=1}^{K} \sum_{j=1}^{k} \lambda_j p_j (1-p_j) \prod_{l\neq j}^{k} p_l$.
 - Using simulations.



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- Why bother?
 - 1. Differences in overall IEO are interesting in their own right.
 - Partial IEOs (looking at the process) and overall IEO (looking at the end result) are natural complements.
 - The fact that partial IEOs do not depend on the distribution of education is both a strong and a weak point.

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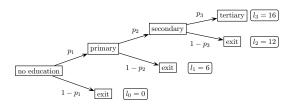
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- ► These are all functions of the transition probabilities.
- These transition probabilities have change over time, differ between men and women, ethnic groups, countries, etc.
- So, differences in overall IEO are in part caused by differences in transition probabilities.



Example

Figure: Hypothetical educational system



Modeling transition probabilities and the expected level of education

$$p_{ki} = \frac{\exp(\alpha_k + \lambda_k SES_i)}{1 + \exp(\alpha_k + \lambda_k SES_i)} \quad \text{if} \quad y_{k-1i} = 1$$

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$$E(ed) = (1 - p_{1i})l_0 + p_{1i}(1 - p_{2i})l_1 + p_{1i}p_{2i}(1 - p_{3i})l_2 + p_{1i}p_{2i}p_{3i}l_3$$

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$$\frac{\partial E(\theta)}{\partial SES} = \{1 \times p_{1i}(1 - p_{1i}) \times [(I_1 - I_0) + p_{2i}(I_2 - I_1) + p_{2i}p_{3i}(I_3 - I_2)]\} \lambda_1 + \{p_{1i} \times p_{2i}(1 - p_{2i}) \times [(I_2 - I_1) + p_{3i}(I_3 - I_2)]\} \lambda_2 + \{p_{1i}p_{2i} \times p_{3i}(1 - p_{3i}) \times [(I_3 - I_2)]\} \lambda_3$$

proportion at risk

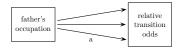
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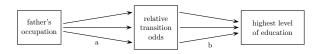
variance of the variable indicating whether one passes or not

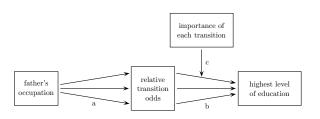
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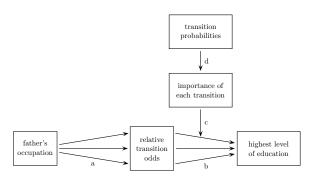
expected increase in the level of education after passing

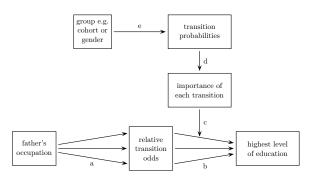
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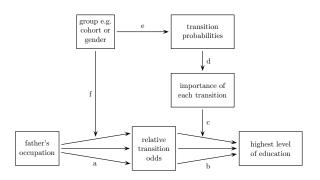












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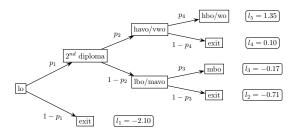
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Simplified model of Dutch educational system



The expected level of education is now:

$$E(ed) = (1 - p_{1i})l_1 + p_{1i}(1 - p_{2i})(1 - p_{3i})l_2 + p_{1i}(1 - p_{2i})p_{3i}l_3 + p_{1i}p_{2i}(1 - p_{4i})l_4 + p_{1i}p_{2i}p_{4i}l_5$$

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The overall IEO is now:

$$\begin{array}{l} \frac{\partial E(ed)}{\partial SES} = \\ \{1 \times p_{1i}(1-p_{1i}) \times & [(1-p_{2i})(l_2-l_1)+p_{2i}(l_4-l_1) + \\ (1-p_{2i})p_{3i}(l_3-l_2)+p_{2i}p_{4i}(l_5-l_4)]\}\lambda_1 + \\ \{p_{1i} \times p_{2i}(1-p_{2i}) \times & [(l_4-l_1)+p_{4i}(l_5-l_4) - \\ (l_2-l_1)-p_{3i}(l_3-l_2)]\}\lambda_2 + \\ \{p_{1i}(1-p_{2i}) \times p_{3i}(1-p_{3i}) \times & [(l_3-l_2)]\}\lambda_3 + \\ \{p_{1i}p_{2i} \times p_{4i}(1-p_{4i}) \times & [(l_5-l_4)]\}\lambda_4 \end{array}$$

The proportion at risk

$$\begin{array}{l} \frac{\partial E(ed)}{\partial SES} = \\ \{\mathbf{1} \times p_{1i}(1-p_{1i}) \times & [(1-p_{2i})(l_2-l_1)+p_{2i}(l_4-l_1)+\\ (1-p_{2i})p_{3i}(l_3-l_2)+p_{2i}p_{4i}(l_5-l_4)]\}\lambda_1 + \\ \{\boldsymbol{p_{1i}} \times p_{2i}(1-p_{2i}) \times & [(l_4-l_1)+p_{4i}(l_5-l_4)-\\ (l_2-l_1)-p_{3i}(l_3-l_2)]\}\lambda_2 + \\ \{\boldsymbol{p_{1i}}(\mathbf{1}-\boldsymbol{p_{2i}}) \times p_{3i}(1-p_{3i}) \times & [(l_3-l_2)]\}\lambda_3 + \\ \{\boldsymbol{p_{1i}}p_{2i} \times p_{4i}(1-p_{4i}) \times & [(l_5-l_4)]\}\lambda_4 \end{array}$$

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- the weights are still a product of
 - the proportion at risk,
 - how far away the pass rate is from universal passing or failing, and
 - the expected increase in level of education from passing
- This result can be generalized to
 - multiple branching points, or
 - branching points with more than two categories.

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- ▶ 51 surveys held between 1958 and 2005 with information on cohorts 1906-1990.
- 67,000 respondents aged between 27 and 65 have complete information on father's occupation, child's education, year of birth, and gender.
- Number of cases are unequally distributed over cohorts.

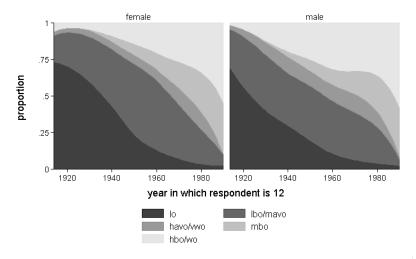
► Father's occupational status is measured in ISEI scores, but recoded to range between 0 and 1.

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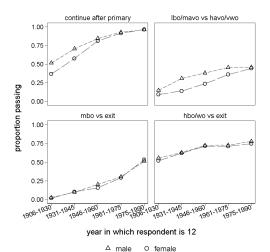
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- Level of education is scaled such as to maximize the direct effect of education on income, and
- it is standardized.
- ► Five cohorts: 1906–1930, 1931–1945, 1946–1960, 1961–1975, 1975–1990.

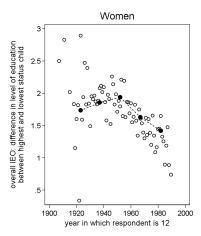
Distribution of highest achieved level of education

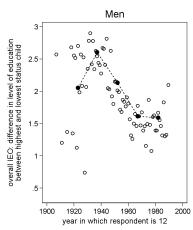


Proportions passing transitions

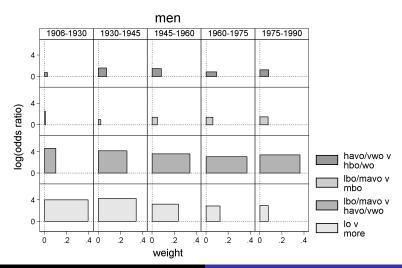


The pattern that needs to be explained

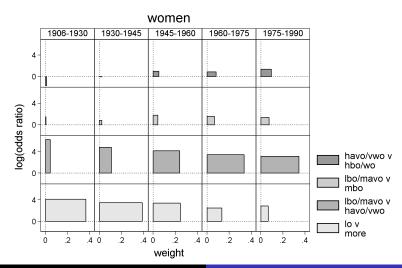




Partial IEOs and their weights for men



Partial IEOs and their weights for women



The first two transitions matter.

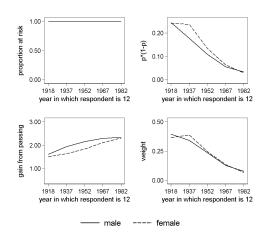
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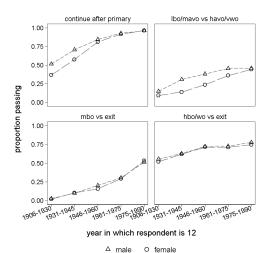
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- The first two transitions matter.
- Overall IEO initially increased because the contribution of the second transition increased.
- Overall IEO later decreased because the contribution of the first transition decreased.
- ▶ This happened more slowly for women, and
- for women the increase and decrease partially canceled each other out.

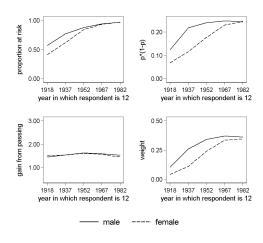
Weight for transition continue after primary



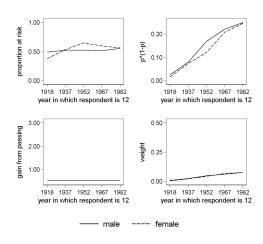
Proportions passing transitions



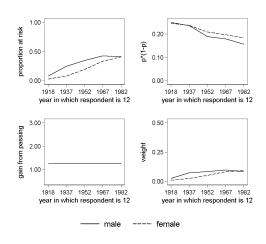
Weight for transition high track v. low track



Weight for transition continue in low track



Weight for transition continue in high track



differences in weights and differences in transition proportions

- Differences between men and women were primarily caused by:
 - the later increase in the proportion of women that continued after primary, and
 - the later increase in the proportion of women that went to the higher track.

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- Overall IEO is a weighted sum of partial IEOs, and the weights increase if:
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 - the proportion that passes is closer to .50,
 - the expected increase in level of education increases

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- This relationship can be used to:
 - identify important and less important transitions,
 - Partial and overall IEO can be related to one another
 - to explain differences in overall IEO with well documented phenomena like educational expansion or the decreased disadvantaged position of women in education.

References



Robert D. Mare.
Change and Stability in Educational Stratification.

American Sociological Review, 46(1):72–87, 1981.

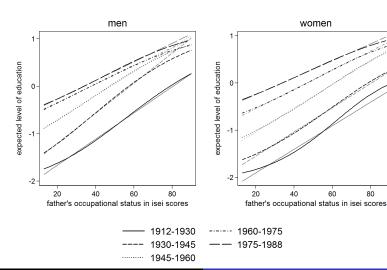
levels of education

English name	before 1968	after 1968	years†	ISCED
primary	LO	lo	6	1
extended primary	VGLO	-	7	1
junior vocational	LTS / ambachtschool	lbo	10	2C
junior vocational	LHNO / huishoudschool	lbo	10	2C
junior general secondary	ULO / MULO	mavo	9 / 10	2B [‡]
senior secondary vocational	MTS	mbo	14	3C
senior general secondary	MMS	havo	11	3B [‡]
pre-university	HBS	vwo	12	3A [‡]
pre-university	lyceum	vwo	12	3A
pre-university	gymnasium	vwo	12	3A
higher professional	HTS	hbo	15	5B
university	universiteit	wo	16	5A

[†] Years refer to the situation after 1968 except VGLO.

[‡] These levels were originally intended to be terminal levels of education for most students (so 2C or 3C) but evolved into levels that primarily grant access to subsequent levels of education.

Mare and OLS





$$ln(inc) = \beta_0 + \underbrace{\beta_1}_{0} lo + \beta_2 lbo_mavo + \beta_3 havo_vwo + \beta_4 mbo + \beta_5 hbo_wo + \cdots$$

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$$\textit{ed} = \underbrace{\alpha_1}_{0} \textit{lo} + \alpha_2 \textit{lbo_mavo} + \alpha_3 \textit{havo_vwo} + \alpha_4 \textit{mbo} + \underbrace{\alpha_5}_{1} \textit{hbo_wo}$$

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$$\begin{array}{ll} \ln(\mathit{inc}) & = & \beta_0 + \gamma_1 \mathit{ed} + \cdots \\ & = & \beta_0 + \gamma_1 \underbrace{\left(\alpha_1 \atop 0} \mathit{lo} + \alpha_2 \mathit{lbo_mavo} + \alpha_3 \mathit{havo_vwo} + \alpha_4 \mathit{mbo} + \underbrace{\alpha_5} \mathit{hbo_wo} \right) + \cdots \end{array}$$

$$\gamma_1 = \beta_5$$

$$\alpha_1 = 0$$

$$\alpha_2 = \frac{\beta_2}{\beta_5}$$

$$\alpha_3 = \frac{\beta_3}{\beta_5}$$

$$\alpha_4 = \frac{\beta_4}{\beta_5}$$

$$\alpha_5 = 1$$

		b	se
α			
	lo	0	
	lbo/mavo	.391	.017
	mbo	.562	.023
	havo/vwo	.659	.022
	hbo/wo	1	
γ			
	1958-1975	.060	.050
	1975-1990	166	.025
	1990-2005	.192	.027
	constant	.474	.074
other			
	1958-1975	.865	.034
	1975-1990	.347	.019
	1990-2005	.161	.022
	fisei	.496	.125
	1958-1975Xfisei	077	.086
	1975-1990Xfisei	132	.044
	1990-2005Xfisei	.073	.042
	age	.115	.004
	age2	071	.003
	constant	4.88	.049